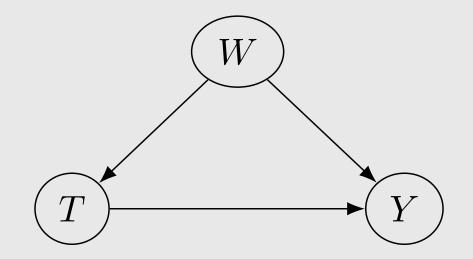
# Unobserved Confounding, Bounds, and Sensitivity Analysis

Brady Neal

causalcourse.com

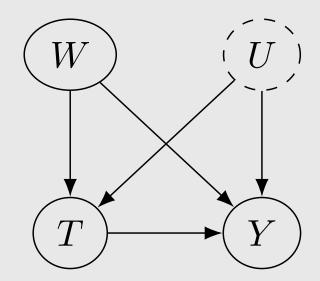
## Motivation: Unobserved Confounding



$$\mathbb{E}[Y(1) - Y(0)] = \mathbb{E}_{W} \left[ \mathbb{E}[Y \mid T = 1, W] - \mathbb{E}[Y \mid T = 0, W] \right]$$

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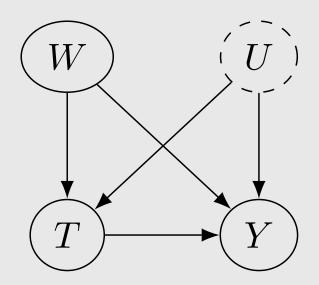
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$$\stackrel{?}{\approx} \mathbb{E}_{W} \left[ \mathbb{E}[Y \mid T = 1, W] - \mathbb{E}[Y \mid T = 0, W] \right]$$

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#### **Bounds**

**No-Assumptions Bound** 

Monotone Treatment Response

**Monotone Treatment Selection** 

**Optimal Treatment Selection** 

#### Sensitivity Analysis

Linear Single Confounder

**Towards More General Settings** 

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Make weaker assumptions:

Identify an interval

"Partial identification" or "set identification"

#### Bounds

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More generally, potential outcomes are bounded:  $\forall t, a \leq Y(t) \leq b$ 

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$$\mathbb{E}[Y(1) - Y(0)]$$

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$$\begin{split} \mathbb{E}[Y(1) - Y(0)] &= \mathbb{E}[Y(1)] - \mathbb{E}[Y(0)] \\ &= P(T=1) \, \mathbb{E}[Y(1) \mid T=1] + P(T=0) \, \mathbb{E}[Y(1) \mid T=0] \\ &- P(T=1) \, \mathbb{E}[Y(0) \mid T=1] - P(T=0) \, \mathbb{E}[Y(0) \mid T=0] \\ &= P(T=1) \, \mathbb{E}[Y \mid T=1] + P(T=0) \, \mathbb{E}[Y(1) \mid T=0] \\ &- P(T=1) \, \mathbb{E}[Y(0) \mid T=1] - P(T=0) \, \mathbb{E}[Y \mid T=0] \end{split}$$

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Observational
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$$\mathbb{E}[Y(1) - Y(0)] = \pi \mathbb{E}[Y \mid T = 1] + (1 - \pi) \mathbb{E}[Y(1) \mid T = 0]$$
 
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## No-Assumptions Bound

Observational-Counterfactual Decomposition

$$\mathbb{E}[Y(1) - Y(0)] = \pi \, \mathbb{E}[Y \mid T = 1] + (1 - \pi) \, \mathbb{E}[Y(1) \mid T = 0]$$
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Recall trivial bound:  $a - b \le \mathbb{E}[Y(1) - Y(0)] \le b - a$ 

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No-assumptions interval length:  $(1 - \pi) b + \pi b - \pi a - (1 - \pi) a$ 

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#### No-Assumptions Bound

$$\mathbb{E}[Y(1) - Y(0)] \le \pi \, \mathbb{E}[Y \mid T = 1] + (1 - \pi) \, b - \pi \, a - (1 - \pi) \, \mathbb{E}[Y \mid T = 0]$$

$$\mathbb{E}[Y(1) - Y(0)] \ge \pi \,\mathbb{E}[Y \mid T = 1] + (1 - \pi) \,a - \pi \,b - (1 - \pi) \,\mathbb{E}[Y \mid T = 0]$$

Plugging these in to the no-assumptions bound:

$$\mathbb{E}[Y(1) - Y(0)] \le (.3)(.9) + (1 - .3)(1) - (.3)(0) - (1 - .3)(.2)$$

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$$-0.17 \le \mathbb{E}[Y(1) - Y(0)] \le 0.83$$

### Questions:

- 1. What kind of bounds can we get on the ATE if the potential outcomes are unbounded?
- 2. Assuming bounded potential outcomes, how much smaller of an interval can we get than the trivial interval [a b, b a]?
- 3. Re-derive the Observational-Counterfactual Decomposition.
- 4. Derive a more general no-assumptions bound where  $a_1 \le Y(1) \le b_1$  and  $a_0 \le Y(0) \le b_0$ .

#### Bounds

No-Assumptions Bound

Monotone Treatment Response

**Monotone Treatment Selection** 

**Optimal Treatment Selection** 

Sensitivity Analysis

Linear Single Confounder

**Towards More General Settings** 

Assume treatment always helps. Mathematically,  $\forall i \ Y_i(1) \geq Y_i(0)$ 

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#### Proof:

$$\mathbb{E}[Y(1) - Y(0)] = \pi \,\mathbb{E}[Y \mid T = 1] + (1 - \pi) \,\mathbb{E}[Y(1) \mid T = 0] - \pi \,\mathbb{E}[Y(0) \mid T = 1] - (1 - \pi) \,\mathbb{E}[Y \mid T = 0]$$

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$$- \pi \, \mathbb{E}[Y(0) \mid T = 1] - (1 - \pi) \, \mathbb{E}[Y \mid T = 0] \qquad \text{Decomposition}$$

$$\geq \pi \, \mathbb{E}[Y \mid T = 1] + (1 - \pi) \, \mathbb{E}[Y \mid T = 0] \qquad \mathbb{E}[Y(1) \mid T = 0]$$

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Potential outcomes bounded between 0 (a) and 1 (b)

$$\pi = 0.3$$

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Combining MTR lower bound with no-assumptions upper bound:

$$0 \le \mathbb{E}[Y(1) - Y(0)] \le 0.83$$

$$\forall i \ Y_i(1) \leq Y_i(0)$$

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#### Running example

No-assumptions bound:  $-0.17 \le \mathbb{E}[Y(1) - Y(0)] \le 0.83$ 

$$\forall i \ Y_i(1) \leq Y_i(0)$$

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#### Running example

No-assumptions bound:  $-0.17 \le \mathbb{E}[Y(1) - Y(0)] \le 0.83$ 

Combining nonpositive MTR upper bound with no-assumptions lower bound:

$$-0.17 \le \mathbb{E}[Y(1) - Y(0)] \le 0$$

### Question:

Given, the nonpositive MTR assumption, prove  $\mathbb{E}[Y(1) - Y(0)] \leq 0$ .

#### Bounds

No-Assumptions Bound

Monotone Treatment Response

**Monotone Treatment Selection** 

**Optimal Treatment Selection** 

#### Sensitivity Analysis

Linear Single Confounder

**Towards More General Settings** 

Treatment groups' potential outcomes are better than control groups':

$$\mathbb{E}[Y(1) \mid T = 1] \ge \mathbb{E}[Y(1) \mid T = 0], \quad \mathbb{E}[Y(0) \mid T = 1] \ge \mathbb{E}[Y(0) \mid T = 0]$$

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Under the MTS assumption, the ATE is bounded from above by the associational difference. Mathematically,

$$\mathbb{E}[Y(1) - Y(0)] \le \mathbb{E}[Y \mid T = 1] - \mathbb{E}[Y \mid T = 0]$$

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Question: Prove the above MTS upper bound.

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No-assumptions bound:  $-0.17 \le \mathbb{E}[Y(1) - Y(0)] \le 0.83$ 

MTS upper bound: 
$$\mathbb{E}[Y(1) - Y(0)] \leq \mathbb{E}[Y \mid T = 1] - \mathbb{E}[Y \mid T = 0]$$

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MTS upper bound: 
$$\mathbb{E}[Y(1) - Y(0)] \leq \mathbb{E}[Y \mid T = 1] - \mathbb{E}[Y \mid T = 0]$$

Combining MTS upper bound with no-assumptions lower bound:

$$-0.17 \le \mathbb{E}[Y(1) - Y(0)] \le 0.7$$

Potential outcomes bounded between 0 (a) and 1 (b)

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  $\mathbb{E}[Y \mid T = 1] = .9$   $\mathbb{E}[Y \mid T = 0] = .2$ 

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$$\mathbb{E}[Y(1) - Y(0)] \leq \mathbb{E}[Y \mid T = 1] - \mathbb{E}[Y \mid T = 0]$$

Combining MTS upper bound with no-assumptions lower bound:  $-0.17 \le \mathbb{E}[Y(1) - Y(0)] \le 0.7$ 

Adding nonnegative MTR assumption and combining MTS upper bound with MTR lower bound ( $\mathbb{E}[Y(1) - Y(0)] \ge 0$ ):

$$0 \le \mathbb{E}[Y(1) - Y(0)] \le 0.7$$

#### Bounds

No-Assumptions Bound

Monotone Treatment Response

**Monotone Treatment Selection** 

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Sensitivity Analysis

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**Towards More General Settings** 

$$T_i = 1 \implies Y_i(1) \ge Y_i(0), \quad T_i = 0 \implies Y_i(0) > Y_i(1)$$

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OTS assumption tells us that  $\mathbb{E}[Y(1) \mid T=0] \leq \mathbb{E}[Y \mid T=0]$ 

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(Observational-Counterfactual Decomposition)

OTS assumption tells us that  $\mathbb{E}[Y(1) \mid T=0] \leq \mathbb{E}[Y \mid T=0]$ 

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$$- \pi \,\mathbb{E}[Y(0) \mid T = 1] - (1 - \pi) \,\mathbb{E}[Y \mid T = 0]$$
$$\leq \pi \,\mathbb{E}[Y \mid T = 1] + (1 - \pi) \,\mathbb{E}[Y \mid T = 0]$$
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$$\leq \pi \, \mathbb{E}[Y \mid T = 1] + (1 - \pi) \, \mathbb{E}[Y \mid T = 0]$$

$$- \pi \, a - (1 - \pi) \, \mathbb{E}[Y \mid T = 0]$$

$$= \pi \, \mathbb{E}[Y \mid T = 1] - \pi \, a$$

OTS assumption tells us that  $\mathbb{E}[Y(0) \mid T=1] \leq \mathbb{E}[Y \mid T=1]$ 

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$$- \pi \, \mathbb{E}[Y \mid T = 1] - (1 - \pi) \, \mathbb{E}[Y \mid T = 0]$$

$$= (1 - \pi) \, a - (1 - \pi) \, \mathbb{E}[Y \mid T = 0]$$

$$\mathbb{E}[Y(1) - Y(0)] < \pi \, \mathbb{E}[Y \mid T = 1] - \pi \, a$$

$$\mathbb{E}[Y(1) - Y(0)] \ge (1 - \pi) \, a - (1 - \pi) \, \mathbb{E}[Y \mid T = 0]$$
Interval Length =  $\pi \, \mathbb{E}[Y \mid T = 1] + (1 - \pi) \, \mathbb{E}[Y \mid T = 0] - a$ 

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OTS Bound 1: 
$$-0.14 \le \mathbb{E}[Y(1) - Y(0)] \le 0.27$$

Interval Length = 0.41

# Bound that identifies the sign

OTS Assumption:  $T_i = 1 \implies Y_i(1) \ge Y_i(0)$ ,  $T_i = 0 \implies Y_i(0) > Y_i(1)$ 

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OTS Bound 1 implication we used:  $\mathbb{E}[Y(1) \mid T=0] \leq \mathbb{E}[Y \mid T=0]$ 

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OTS Bound 2 implication we'll use:  $\mathbb{E}[Y(1) \mid T=0] \leq \mathbb{E}[Y \mid T=1]$ 

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Proof: 
$$\mathbb{E}[Y(1) \mid T = 0] = \mathbb{E}[Y(1) \mid Y(0) > Y(1)]$$

OTS Assumption:  $T_i = 1 \implies Y_i(1) \ge Y_i(0)$ ,  $T_i = 0 \implies Y_i(0) > Y_i(1)$ 

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$$\mathbb{E}[Y(1) \mid T = 0] = \mathbb{E}[Y(1) \mid Y(0) > Y(1)]$$
  
  $\leq \mathbb{E}[Y(1) \mid Y(0) \leq Y(1)]$ 

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Proof: 
$$\mathbb{E}[Y(1) \mid T = 0] = \mathbb{E}[Y(1) \mid Y(0) > Y(1)]$$
  
 $\leq \mathbb{E}[Y(1) \mid Y(0) \leq Y(1)]$   
 $= \mathbb{E}[Y(1) \mid T = 1]$ 

OTS Assumption:  $T_i = 1 \implies Y_i(1) \ge Y_i(0)$ ,  $T_i = 0 \implies Y_i(0) > Y_i(1)$ Contrapositive:  $T_i = 1 \iff Y_i(0) \le Y_i(1)$ 

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 $= \mathbb{E}[Y(1) \mid T = 1]$   
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OTS assumption tells us that  $\mathbb{E}[Y(1) \mid T=0] \leq \mathbb{E}[Y \mid T=1]$ 

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$$- \pi \,\mathbb{E}[Y(0) \mid T = 1] - (1 - \pi) \,\mathbb{E}[Y \mid T = 0]$$

$$\leq \pi \,\mathbb{E}[Y \mid T = 1] + (1 - \pi) \,\mathbb{E}[Y \mid T = 1]$$

$$- \pi \,a - (1 - \pi) \,\mathbb{E}[Y \mid T = 0]$$

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$$- \pi \, \mathbb{E}[Y(0) \mid T = 1] - (1 - \pi) \, \mathbb{E}[Y \mid T = 0]$$

$$\leq \pi \, \mathbb{E}[Y \mid T = 1] + (1 - \pi) \, \mathbb{E}[Y \mid T = 1]$$

$$- \pi \, a - (1 - \pi) \, \mathbb{E}[Y \mid T = 0]$$

$$= \mathbb{E}[Y \mid T = 1] - \pi \, a - (1 - \pi) \, \mathbb{E}[Y \mid T = 0]$$

## Question:

Prove a new lower bound using the version of the OTS assumption that we used in the last slide.

$$\mathbb{E}[Y(1) - Y(0)] \le \mathbb{E}[Y \mid T = 1] - \pi a - (1 - \pi) \mathbb{E}[Y \mid T = 0]$$

$$\mathbb{E}[Y(1) - Y(0)] \ge \pi \mathbb{E}[Y \mid T = 1] + (1 - \pi) a - \mathbb{E}[Y \mid T = 0]$$
Interval Length =  $(1 - \pi) \mathbb{E}[Y \mid T = 1] + \pi \mathbb{E}[Y \mid T = 0] - a$ 

$$\mathbb{E}[Y(1) - Y(0)] \le \mathbb{E}[Y \mid T = 1] - \pi \, a - (1 - \pi) \, \mathbb{E}[Y \mid T = 0]$$

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$$0.07 \le \mathbb{E}[Y(1) - Y(0)] \le 0.76$$
  
Interval Length = 0.69

$$\mathbb{E}[Y(1) - Y(0)] \le \mathbb{E}[Y \mid T = 1] - \pi a - (1 - \pi) \mathbb{E}[Y \mid T = 0]$$

$$\mathbb{E}[Y(1) - Y(0)] \ge \pi \mathbb{E}[Y \mid T = 1] + (1 - \pi) a - \mathbb{E}[Y \mid T = 0]$$
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Identified the sign of the effect!

# Comparing and Mixing OTS Bounds

Potential outcomes bounded between 0 (a) and 1 (b)

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# Comparing and Mixing OTS Bounds

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Identified the sign of the effect, but gives a 68% larger interval

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Interval Length = 0.69

Identified the sign of the effect, but gives a 68% larger interval

OTS Upper Bound 1 and OTS Lower Bound 2:

$$0.07 \le \mathbb{E}[Y(1) - Y(0)] \le 0.27$$

Interval Length = 0.2

#### Bounds

No-Assumptions Bound

Monotone Treatment Response

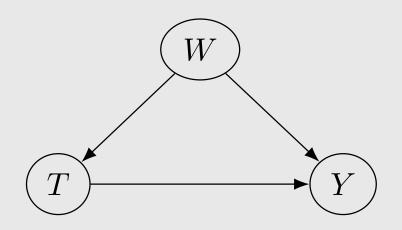
**Monotone Treatment Selection** 

**Optimal Treatment Selection** 

#### Sensitivity Analysis

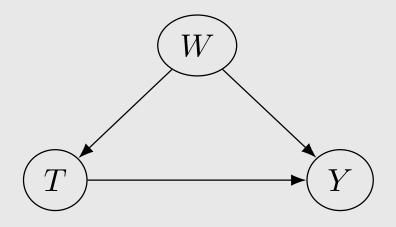
Linear Single Confounder

**Towards More General Settings** 



$$\mathbb{E}[Y(1) - Y(0)] = \mathbb{E}_{W} \left[ \mathbb{E}[Y \mid T = 1, W] - \mathbb{E}[Y \mid T = 0, W] \right]$$

Last section, we completely threw out the unconfoundedness assumption.

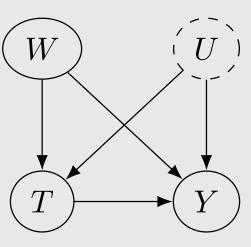


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Now, we assume the observed W and the unobserved U gives us

unconfoundedness

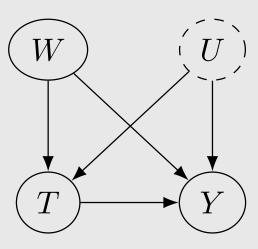


$$\mathbb{E}[Y(1) - Y(0)] = \mathbb{E}_{W,U} \left[ \mathbb{E}[Y \mid T = 1, W, U] - \mathbb{E}[Y \mid T = 0, W, U] \right]$$

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$$\stackrel{?}{\approx} \mathbb{E}_{W} \left[ \mathbb{E}[Y \mid T = 1, W] - \mathbb{E}[Y \mid T = 0, W] \right]$$

#### **Bounds**

No-Assumptions Bound

Monotone Treatment Response

**Monotone Treatment Selection** 

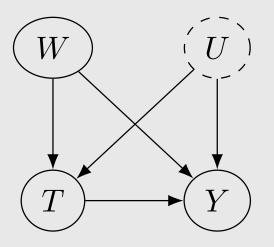
**Optimal Treatment Selection** 

#### Sensitivity Analysis

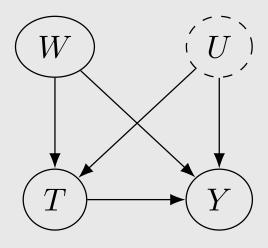
Linear Single Confounder

**Towards More General Settings** 

#### Linear SCM



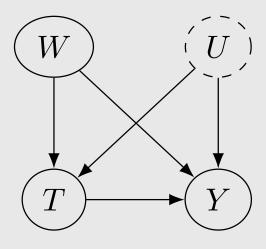
#### Linear SCM



$$T := \alpha_w W + \alpha_u U$$

$$Y := \beta_w W + \beta_u U + \delta T$$

#### Linear SCM



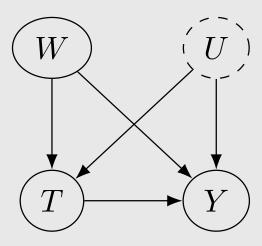
$$T := \alpha_w W + \alpha_u U$$

$$Y := \beta_w W + \beta_u U + \underline{\delta} T$$

Goal: recover  $\delta$ 

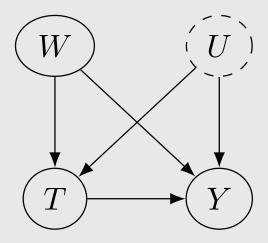
$$T := \alpha_w W + \alpha_u U$$

$$Y := \beta_w W + \beta_u U + \delta T$$



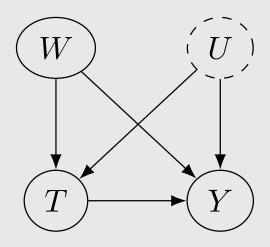
$$T := \alpha_w W + \alpha_u U$$

$$Y := \beta_w W + \beta_u U + \delta T$$



$$\mathbb{E}[Y(1) - Y(0)] = \mathbb{E}_{W,U} \left[ \mathbb{E}[Y \mid T = 1, W, U] - \mathbb{E}[Y \mid T = 0, W, U] \right] = \delta$$

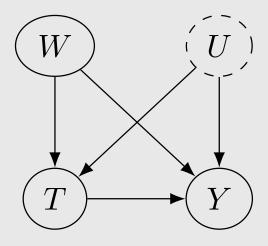
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$$\mathbb{E}_{W} \left[ \mathbb{E}[Y \mid T = 1, W] - \mathbb{E}[Y \mid T = 0, W] \right] \stackrel{?}{=}$$

$$T := \alpha_w W + \alpha_u U$$
$$Y := \beta_w W + \beta_u U + \delta T$$

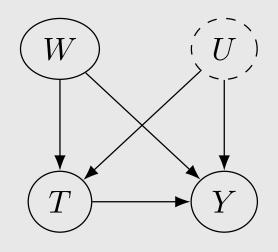


$$\mathbb{E}[Y(1) - Y(0)] = \mathbb{E}_{W,U} \left[ \mathbb{E}[Y \mid T = 1, W, U] - \mathbb{E}[Y \mid T = 0, W, U] \right] = \delta$$

$$\mathbb{E}_{W} \left[ \mathbb{E}[Y \mid T = 1, W] - \mathbb{E}[Y \mid T = 0, W] \right] = \delta + \frac{\beta_{u}}{\alpha_{u}}$$

$$T := \alpha_w W + \alpha_u U$$

$$Y := \beta_w W + \beta_u U + \delta T$$



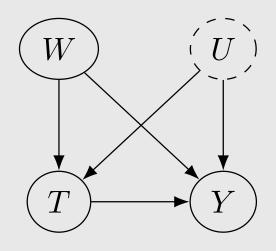
Proof coming after next slide

$$\mathbb{E}[Y(1) - Y(0)] = \mathbb{E}_{W,U} \left[ \mathbb{E}[Y \mid T = 1, W, U] - \mathbb{E}[Y \mid T = 0, W, U] \right] = \delta$$

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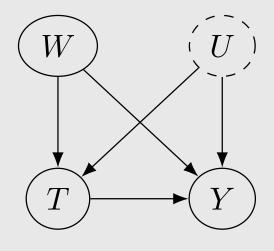
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Bias of 
$$\mathbb{E}_W \left[ \mathbb{E}[Y \mid T = 1, W] - \mathbb{E}[Y \mid T = 0, W] \right] = \delta + \frac{\beta_u}{\alpha_u} - \delta$$

$$T := \alpha_w W + \alpha_u U$$
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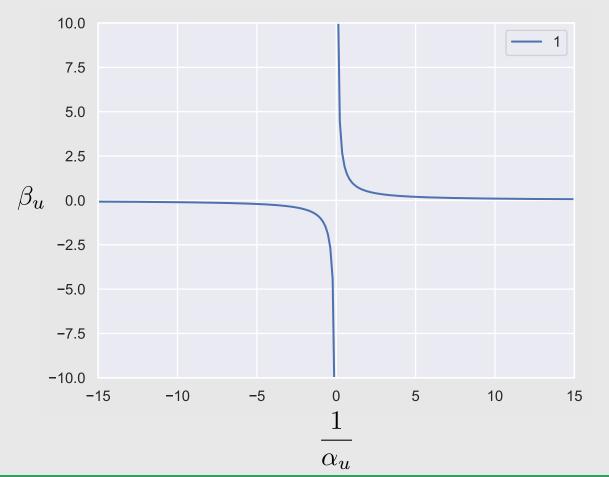


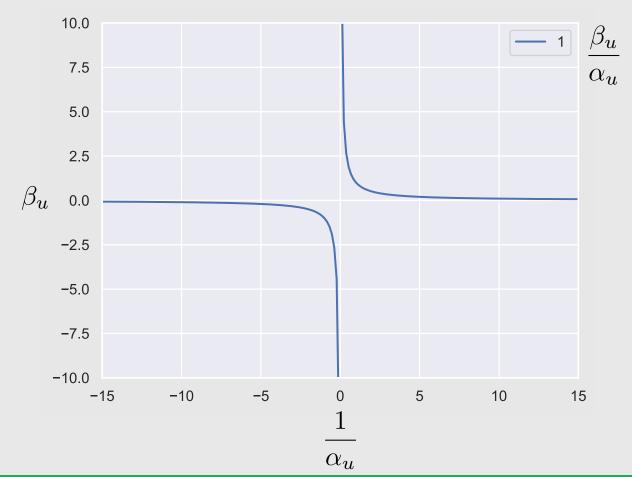
Proof coming after next slide

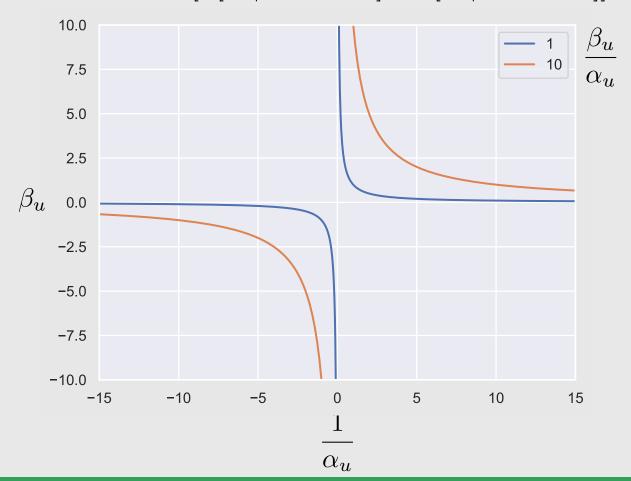
$$\mathbb{E}[Y(1) - Y(0)] = \mathbb{E}_{W,U} \left[ \mathbb{E}[Y \mid T = 1, W, U] - \mathbb{E}[Y \mid T = 0, W, U] \right] = \delta$$

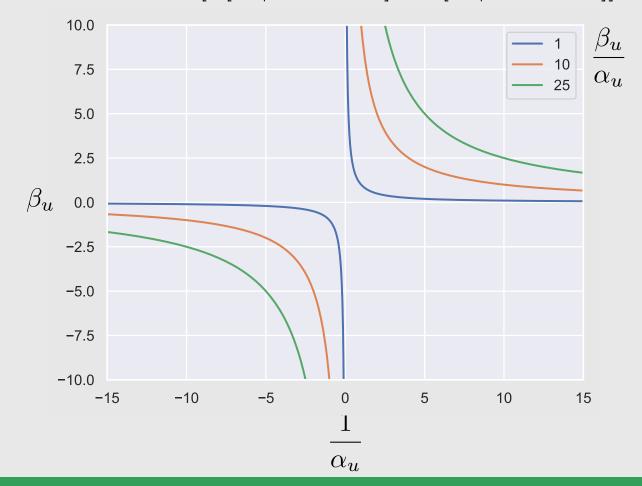
$$\mathbb{E}_{W} \left[ \mathbb{E}[Y \mid T = 1, W] - \mathbb{E}[Y \mid T = 0, W] \right] = \delta + \frac{\beta_{u}}{\alpha_{u}}$$

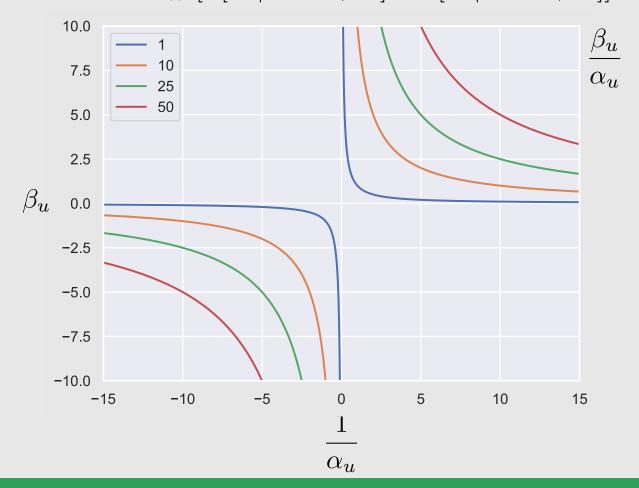
Bias of 
$$\mathbb{E}_W \left[ \mathbb{E}[Y \mid T = 1, W] - \mathbb{E}[Y \mid T = 0, W] \right] = \delta + \frac{\beta_u}{\alpha_u} - \delta = \frac{\beta_u}{\alpha_u}$$

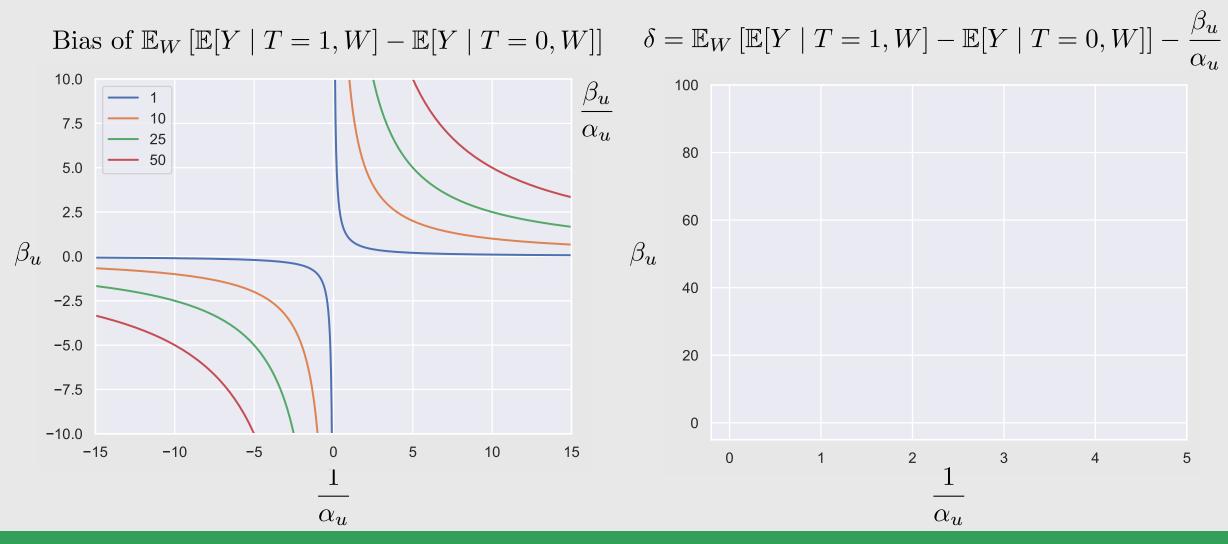


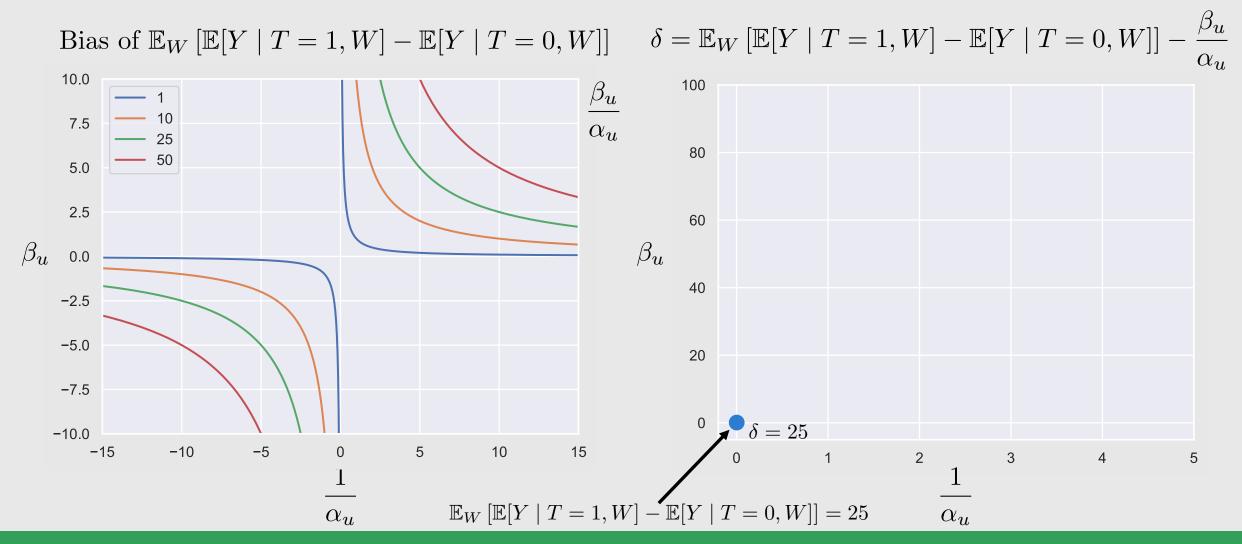


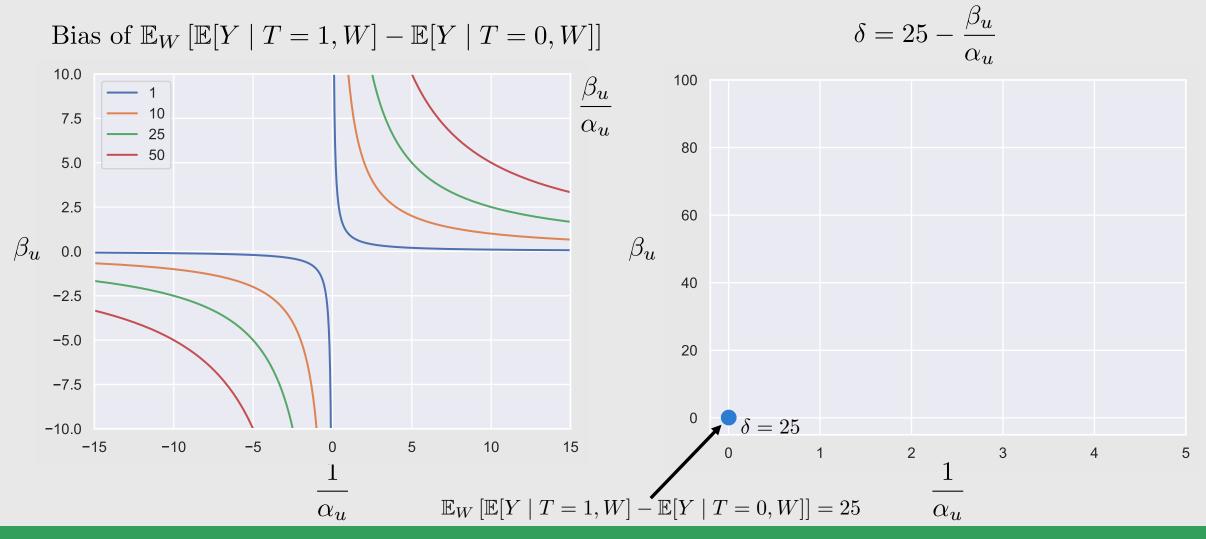


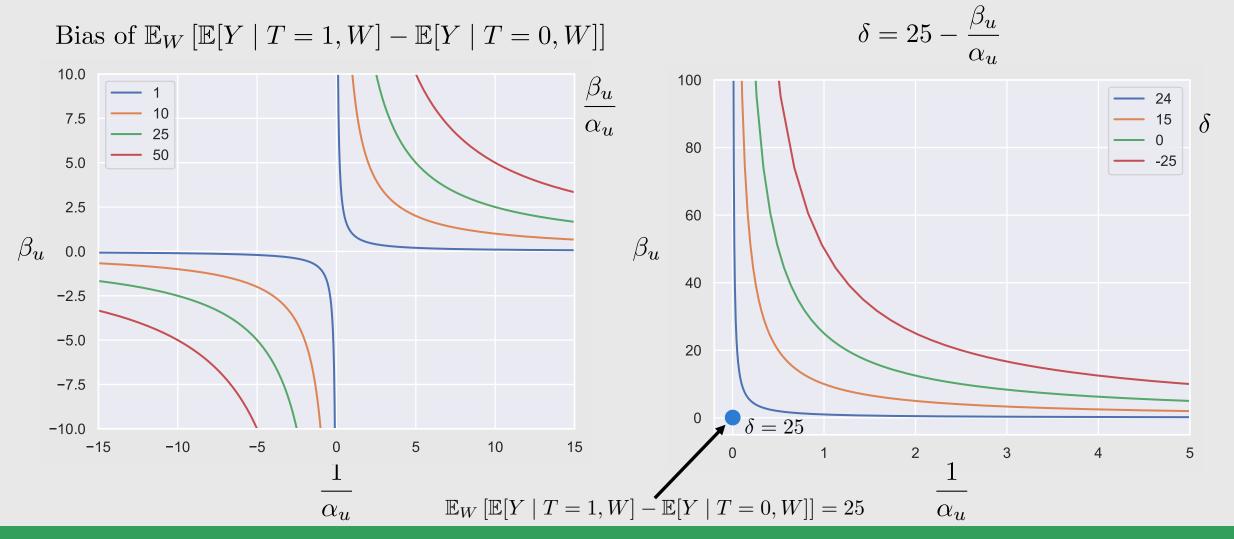












Assumed SCM:

$$T := \alpha_w W + \alpha_u U$$

$$Y := \beta_w W + \beta_u U + \delta T$$

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$$T := \alpha_w W + \alpha_u U$$
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Result:

The confounding bias of adjusting for just W (and not U) is  $\frac{\beta_u}{\alpha_u}$ . Formally,

$$\mathbb{E}_{W} \left[ \mathbb{E}[Y \mid T = 1, W] - \mathbb{E}[Y \mid T = 0, W] \right]$$
$$- \mathbb{E}_{W,U} \left[ \mathbb{E}[Y \mid T = 1, W, U] - \mathbb{E}[Y \mid T = 0, W, U] \right] = \frac{\beta_{u}}{\alpha_{u}}$$

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Proof Outline:

1. Get a closed-form expression for  $\mathbb{E}_W [\mathbb{E}[Y \mid T = t, W]]$  in terms of  $\alpha_w$ ,  $\alpha_u$ ,  $\beta_w$ , and  $\beta_u$ .

Assumed SCM: 
$$T := \alpha_w W + \alpha_u U$$
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$$T := \alpha_w W + \alpha_u U$$
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- 1. Get a closed-form expression for  $\mathbb{E}_W [\mathbb{E}[Y \mid T = t, W]]$  in terms of  $\alpha_w$ ,  $\alpha_u$ ,  $\beta_w$ , and  $\beta_u$ .
- 2. Use step 1 to get a closed-form expression for the difference  $\mathbb{E}_W \left[ \mathbb{E}[Y \mid T=1, W] \mathbb{E}[Y \mid T=0, W] \right]$
- 3. Subtract off  $\mathbb{E}_{W,U} [\mathbb{E}[Y \mid T=1, W, U] \mathbb{E}[Y \mid T=0, W, U]] = \delta$

Assumed SCM: 
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Assumed SCM: 
$$T := \alpha_w W + \alpha_u U$$
$$Y := \beta_w W + \beta_u U + \delta T$$

$$\mathbb{E}_W \left[ \mathbb{E}[Y \mid T = t, W] \right]$$

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$$T := \alpha_w W + \alpha_u U$$

$$Y := \beta_w W + \beta_u U + \delta T$$

$$\mathbb{E}_W \left[ \mathbb{E}[Y \mid T = t, W] \right]$$

$$T := \alpha_w W + \alpha_u U$$
$$Y := \beta_w W + \beta_u U + \delta T$$

$$\mathbb{E}_{W} \left[ \mathbb{E}[Y \mid T = t, W] \right] = \mathbb{E}_{W} \left[ \mathbb{E}[\beta_{w}W + \beta_{u}U + \delta T \mid T = t, W] \right]$$

Assumed SCM: 
$$T := \alpha_w W + \alpha_u U$$
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$$= \mathbb{E}_{W} \left[ \beta_{w}W + \beta_{u}\mathbb{E}[U \mid T = t, W] + \delta t \right]$$

Assumed SCM: 
$$\frac{T := \alpha_w W + \alpha_u U}{Y := \beta_w W + \beta_u U + \delta T}$$

$$\mathbb{E}_{W} \left[ \mathbb{E}[Y \mid T = t, W] \right] = \mathbb{E}_{W} \left[ \mathbb{E}[\beta_{w}W + \beta_{u}U + \delta T \mid T = t, W] \right]$$
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Assumed SCM: 
$$\frac{T := \alpha_w W + \alpha_u U}{Y := \beta_w W + \beta_u U + \delta T} \qquad U = \frac{T - \alpha_w W}{\alpha_u}$$

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$$= \mathbb{E}_{W} \left[ \beta_{w}W + \beta_{u}\mathbb{E}[U \mid T = t, W] + \delta t \right]$$

$$= \mathbb{E}_{W} \left[ \beta_{w}W + \beta_{u} \left( \frac{t - \alpha_{w}W}{\alpha_{u}} \right) + \delta t \right]$$

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$$= \mathbb{E}_{W} \left[ \beta_{w}W + \frac{\beta_{u}}{\alpha_{u}}t - \frac{\beta_{u}\alpha_{w}}{\alpha_{u}}W + \delta t \right]$$

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$$= \mathbb{E}_{W} \left[ \beta_{w}W + \frac{\beta_{u}}{\alpha_{u}}t - \frac{\beta_{u}\alpha_{w}}{\alpha_{u}}W + \delta t \right]$$

$$= \beta_{w}\mathbb{E}[W] + \frac{\beta_{u}}{\alpha_{u}}t - \frac{\beta_{u}\alpha_{w}}{\alpha_{u}}\mathbb{E}[W] + \delta t$$

Assumed SCM: 
$$T := \alpha_w W + \alpha_u U$$
 
$$Y := \beta_w W + \beta_u U + \delta T$$
 
$$U = \frac{T - \alpha_w W}{\alpha_u}$$

$$\begin{split} \mathbb{E}_{W} \left[ \mathbb{E}[Y \mid T = t, W] \right] &= \mathbb{E}_{W} \left[ \mathbb{E}[\beta_{w}W + \beta_{u}U + \delta T \mid T = t, W] \right] \\ &= \mathbb{E}_{W} \left[ \beta_{w}W + \beta_{u}\mathbb{E}[U \mid T = t, W] + \delta t \right] \\ &= \mathbb{E}_{W} \left[ \beta_{w}W + \beta_{u} \left( \frac{t - \alpha_{w}W}{\alpha_{u}} \right) + \delta t \right] \\ &= \mathbb{E}_{W} \left[ \beta_{w}W + \frac{\beta_{u}}{\alpha_{u}}t - \frac{\beta_{u}\alpha_{w}}{\alpha_{u}}W + \delta t \right] \\ &= \beta_{w}\mathbb{E}[W] + \frac{\beta_{u}}{\alpha_{u}}t - \frac{\beta_{u}\alpha_{w}}{\alpha_{u}}\mathbb{E}[W] + \delta t \\ &= \left( \delta + \frac{\beta_{u}}{\alpha_{u}} \right)t + \left( \beta_{w} - \frac{\beta_{u}\alpha_{w}}{\alpha_{u}} \right)\mathbb{E}[W] \end{split}$$

Step 1: 
$$\mathbb{E}_W \left[ \mathbb{E}[Y \mid T = t, W] \right] = \left( \delta + \frac{\beta_u}{\alpha_u} \right) t + \left( \beta_w - \frac{\beta_u \alpha_w}{\alpha_u} \right) \mathbb{E}[W]$$

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$$\mathbb{E}_{W} \left[ \mathbb{E}[Y \mid T = 1, W] - \mathbb{E}[Y \mid T = 0, W] \right]$$

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$$\mathbb{E}_{W} \left[ \mathbb{E}[Y \mid T = 1, W] - \mathbb{E}[Y \mid T = 0, W] \right] = \left( \delta + \frac{\beta_{u}}{\alpha_{u}} \right) (1) + \left( \beta_{w} - \frac{\beta_{u} \alpha_{w}}{\alpha_{u}} \right) \mathbb{E}[W]$$

$$- \left[ \left( \delta + \frac{\beta_{u}}{\alpha_{u}} \right) (0) + \left( \beta_{w} - \frac{\beta_{u} \alpha_{w}}{\alpha_{u}} \right) \mathbb{E}[W] \right]$$

Brady Neal

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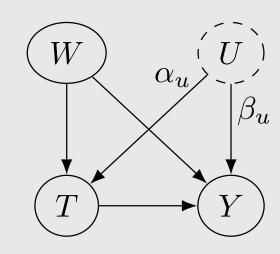
$$= \delta + \frac{\beta_{u}}{\alpha_{u}}$$

Bias = 
$$\mathbb{E}_{W} [\mathbb{E}[Y \mid T = 1, W] - \mathbb{E}[Y \mid T = 0, W]]$$
  
-  $\mathbb{E}_{W,U} [\mathbb{E}[Y \mid T = 1, W, U] - \mathbb{E}[Y \mid T = 0, W, U]]$ 

$$\begin{aligned} \operatorname{Bias} &= \mathbb{E}_{W} \left[ \mathbb{E}[Y \mid T = 1, W] - \mathbb{E}[Y \mid T = 0, W] \right] \\ &- \mathbb{E}_{W,U} \left[ \mathbb{E}[Y \mid T = 1, W, U] - \mathbb{E}[Y \mid T = 0, W, U] \right] \\ &= \delta + \frac{\beta_{u}}{\alpha_{u}} - \delta \end{aligned}$$

Bias = 
$$\mathbb{E}_{W} [\mathbb{E}[Y \mid T = 1, W] - \mathbb{E}[Y \mid T = 0, W]]$$
  
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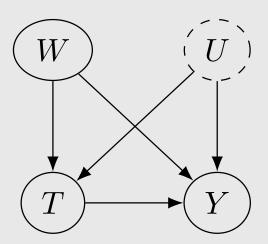


$$T := \alpha_w W + \alpha_u U$$

$$Y := \beta_w W + \beta_u U + \delta T$$

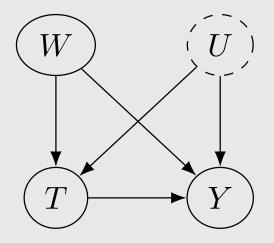
# Generalization to Arbitrary Linear SCMs

We've considered specifically the ATE in this simple graph  $\mathbb{E}[Y(1)-Y(0)]$ 



### Generalization to Arbitrary Linear SCMs

We've considered specifically the ATE in this simple graph  $\mathbb{E}[Y(1)-Y(0)]$ 



See "Sensitivity Analysis of Linear Structural Causal Models" from Cinelli et al. (2019) for arbitrary estimands in arbitrary graphs, where the structural equations are still linear

SCM: 
$$T := \alpha_w W + \alpha_u U$$
$$Y := \beta_w W + \beta_u U + \delta T$$

#### Questions:

- 1. Given the above SCM, show that  $\mathbb{E}_{W,U}\left[\mathbb{E}[Y\mid T=1,W,U]-\mathbb{E}[Y\mid T=0,W,U]\right]=\delta$
- 2. Does what we have shown in this section work if W is a vector?
- 3. How about if U is a vector?

#### Bounds

No-Assumptions Bound

Monotone Treatment Response

**Monotone Treatment Selection** 

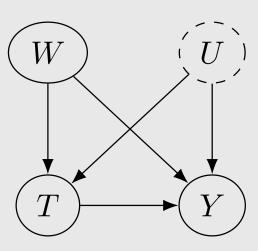
**Optimal Treatment Selection** 

Sensitivity Analysis

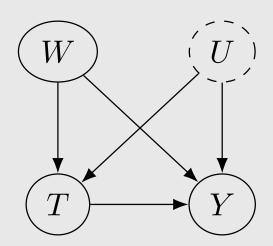
Linear Single Confounder

**Towards More General Settings** 

$$T := \alpha_w W + \alpha_u U$$
$$Y := \beta_w W + \beta_u U + \delta T$$



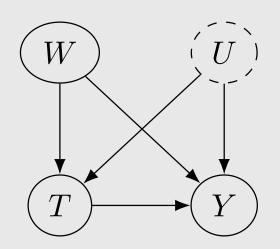
$$T := \alpha_w W + \alpha_u U$$
$$Y := \beta_w W + \beta_u U + \delta T$$



$$P(T = 1 \mid W, U) := \text{sigmoid} (\alpha_w W + \alpha_u U)$$

$$Y := \beta_w W + \beta_u U + \delta T + N$$
where  $\text{sigmoid}(x) = \frac{1}{1 + e^{-x}}$ 

$$T := \alpha_w W + \alpha_u U$$
$$Y := \beta_w W + \beta_u U + \delta T$$



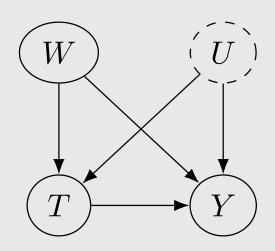
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where sigmoid(x) =  $\frac{1}{1 + e^{-x}}$ 

Rosenbaum & Rubin (1983) and Imbens (2003)

• Simple parametric form for T

$$T := \alpha_w W + \alpha_u U$$
$$Y := \beta_w W + \beta_u U + \delta T$$

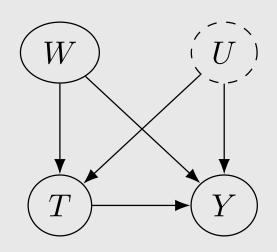


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- Simple parametric form for T
- Simple parametric form for Y

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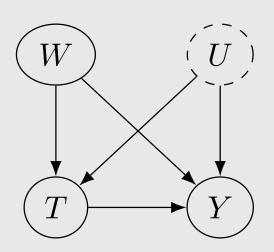


$$P(T = 1 \mid W, U) := \text{sigmoid} (\alpha_w W + \alpha_u U)$$

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- Simple parametric form for T
- Simple parametric form for Y
- U is binary

$$T := \alpha_w W + \alpha_u U$$
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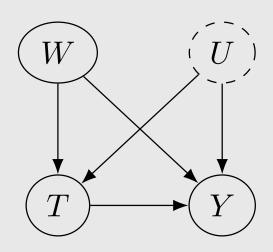


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- Simple parametric form for T
- Simple parametric form for Y
- U is binary
- U is a scalar (only one unobserved confounder)

$$T := \alpha_w W + \alpha_u U$$
$$Y := \beta_w W + \beta_u U + \delta T$$



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$$Y := \beta_w W + \beta_u U + \delta T + N$$
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Cinelli & Hazlett (2020) drop many of these assumptions

- Simple parametric form for T
- Simple parametric form for Y
- U is binary
- U is a scalar (only one unobserved confounder)

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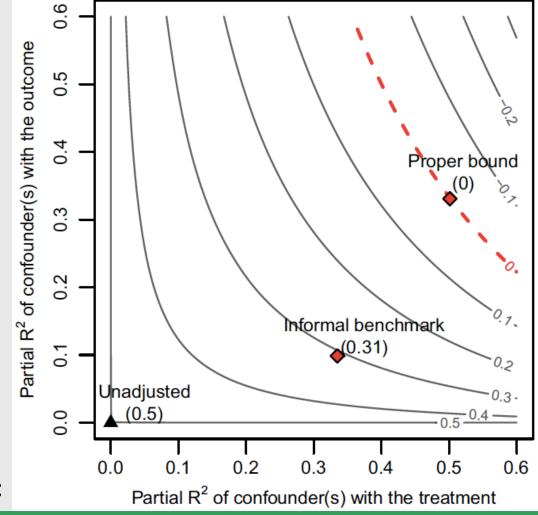


Figure 4:

- Assumes no parametric form for structural equation for T
- Allows for multiple confounders (U can be a vector)
- No assumed distribution on U
- Gives rigorous bound to tells us when we can know that our causal effect estimate is sufficiently robust to and follow-ups unobserved confounding

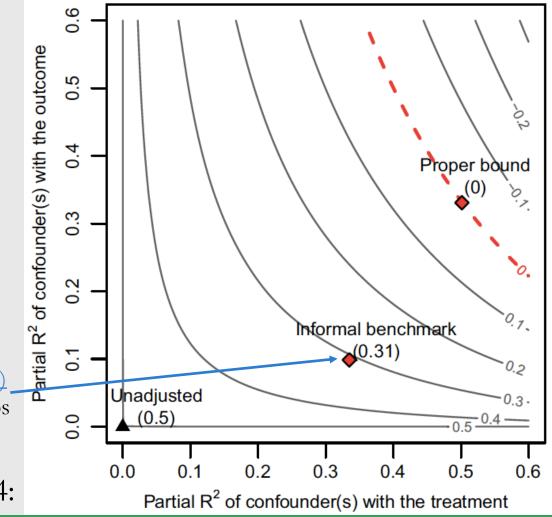


Figure 4:

Sense and Sensitivity Analysis Veitch & Zaveri (2020)

#### Sense and Sensitivity Analysis Veitch & Zaveri (2020)

Gives the same nice sensitivity curves we've been seeing, but with more flexible parameterization

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Gives the same nice sensitivity curves we've been seeing, but with more flexible parameterization

Both the treatment mechanism and the outcome mechanism can be modeled with **arbitrary machine learning models**, and we still get a closed-form expression for the bias (assuming well-specification)

### Lots of Other Sensitivity Analysis Methods

- An Introduction to Sensitivity Analysis for Unobserved Confounding in Non-Experimental Prevention Research (Liu, Kuramoto, & Stuart., 2013)
- Rosenbaum has several (Rosenbaum 2002, 2010, 2017)
- Unmeasured Confounding for General Outcomes, Treatments, and Confounders (VanderWeele & Arah, 2011)
- Sensitivity Analysis Without Assumptions (Ding & VanderWeele, 2018)
- Flexible sensitivity analysis for observational studies without observable implications (Franks, D'Amour, & Feller, 2019)
- Bounds on the conditional and average treatment effect with unobserved confounding factors (Yadlowsky et al., 2018)